

Large Mixed-Frequency VARs with a Parsimonious Time-Varying Parameter Structure

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Introduction

To simultaneously consider mixed-frequency time series, their joint dynamics, and possible structural changes, we introduce a time-varying parameter mixed-frequency VAR (MF-VAR). To keep our approach from becoming too complex, we **implement time variation parsimoniously**: only the intercepts and a common factor in the error variances vary over time. We can therefore estimate moderately large systems in a reasonable amount of time making our modifications appealing for practical use. We further improve upon **computational efficiency** by exploiting gains from using sparse and block-banded matrix algebra as, for instance, in Chan (2015). Furthermore, we complement the standard Minnesota prior setting to long-run priors as in Giannone et al. (2016). For eleven U.S. variables, we examine the performance of our model (and two intermediate variants), **first and foremost with respect to GDP forecasting**, and compare them to the time-constant MF-VAR of Schorfheide and Song (2015). Our results demonstrate the feasibility and usefulness of our method.

Motivating Example – Restricting Time Variation

- Our model lies “in-between” the time-constant MF-VAR of Schorfheide and Song (2015) and the fully fledged time-varying model of Cimadomo and D’Agostino (2015), thereby handling the trade-off between model flexibility and complexity
- For an $n = 3$ -dimensional (GDP, CPI, fed funds rate) system we compare our model with a fully time-varying model $\Rightarrow n + 1$ instead of $n + n^2p + \frac{n(n+1)}{2}$ time-varying quantities (p - lag length)

Table: Forecast Performance relative to a MF-VAR à la Primiceri (2005)

Horizon	-1	0	1	2	3	4	5	6	7	8	9	10
Rel. RMSE	1.02	0.94	0.84	1.01	0.95	0.86	0.95	0.91	0.86	0.92	0.89	0.86
Rel. logPL	-0.64	-0.19	0.04	0.19	0.18	0.23	0.18	0.25	0.14	0.11	0.16	-0.04

Note: -1 is a backcast, 0 and 1 are nowcasts, the remaining horizons refer to forecasts. More on forecast horizons later.

- Rarely worse performance than the fully fledged MF-VAR \Rightarrow time variation in the data seems to be sufficiently covered by intercepts and (common) stochastic volatility

Table: Running Time relative to a MF-VAR à la Primiceri (2005)

p	n		
	3	4	6
3	42.9	34.4	10.6
6	33.9	16.4	
12	25.0		

- Blank entries indicate the unrestricted model to run into computational problems \Rightarrow moderately large systems like we have in mind ($n = 10$ to 15) become practically infeasible to run

Our Modeling Framework – Two Blocks

- Distinguish between a monthly and a “quarterly” part:

$$\begin{bmatrix} x_{m,t} \\ x_{q,t} \end{bmatrix} = \begin{bmatrix} \Phi_{mm} & \Phi_{mq} & \Phi_{mc} \\ \Phi_{qm} & \Phi_{qq} & \Phi_{qc} \end{bmatrix} \begin{bmatrix} x_{m,t-1} \\ x_{q,t-1} \\ 1 \end{bmatrix} + \begin{bmatrix} \Phi_{mc,t} \\ \Phi_{qc,t} \end{bmatrix} + \begin{bmatrix} u_{m,t} \\ u_{q,t} \end{bmatrix}$$

$$\begin{bmatrix} u_{m,t} \\ u_{q,t} \end{bmatrix} \sim N \left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} \Sigma_{mm,t} & \Sigma_{mq,t} \\ \Sigma_{mq,t} & \Sigma_{qq,t} \end{bmatrix} \right)$$

- $X_{m,t} = (x'_{m,t}, \dots, x'_{m,t-p-1})'$ and $X_{q,t} = (x'_{q,t}, \dots, x'_{q,t-p-1})'$ are $pn_m \times 1$ and $pn_q \times 1$ vectors (potentially latent observations; observable ones denoted $Y_{m,t}$ and $Y_{q,t}$)
- Compactly (i.e., stacked over t), $X = \Phi Z + \Phi_c^T + U$ with $u = \text{vec}(U) \sim N(0, \Omega \otimes \Psi)$, where Ψ governs “usual” cross-sectional and Ω serial structures (Chan, 2015) \Rightarrow this **Kronecker structure** proves quite useful for computational efficiency
- Block I** (dealing with the “incomplete” data given Block II):

* For the balanced part of the sample, “reduced” state vector, $S_t = (x'_{m,t}, \dots, x'_{q,t-p})'$, noting that $Y_{m,t} = X_{m,t}$:

$$S_t = \begin{bmatrix} \Phi_{qq} & 0 \\ I_{pn_q} & 0 \end{bmatrix} S_{t-1} + \begin{bmatrix} \Phi_{qm} \\ 0 \end{bmatrix} Y_{m,t-1} + \begin{bmatrix} \Phi_{qc,t} \\ 0 \end{bmatrix} + \begin{bmatrix} u_{q,t} \\ 0 \end{bmatrix} \text{ (state equation)}$$

$$\begin{bmatrix} y_{m,t} \\ y_{q,t} \end{bmatrix} = \begin{bmatrix} 0 & \Phi_{mq,1} & \Phi_{mq,2} & \Phi_{mq,3,p} \\ 1/3 I_{n_q} & 1/3 I_{n_q} & 1/3 I_{n_q} & 0 \end{bmatrix} S_t + \begin{bmatrix} \Phi_{mm} \\ 0 \end{bmatrix} Y_{m,t-1} + \begin{bmatrix} \Phi_{mc,t} \\ 0 \end{bmatrix} + \begin{bmatrix} u_{m,t} \\ 0 \end{bmatrix} \text{ (measurement equation)}$$

* For the ragged edge (unbalanced) part, $S_t = (x'_{m,t}, \dots, x'_{q,t-p})'$, where $x_t = (x'_{m,t}, x'_{q,t})'$

$$S_t = \begin{bmatrix} \Phi_{mm,1} & \Phi_{mq,1} & \dots & \Phi_{mm,p} & \Phi_{mq,p} & 0 \\ \Phi_{qm,1} & \Phi_{qq,1} & \dots & \Phi_{qm,p} & \Phi_{qq,p} & 0 \end{bmatrix} S_{t-1} + \begin{bmatrix} \Phi_{mc} + \Phi_{mc,t} \\ \Phi_{qc} + \Phi_{qc,t} \end{bmatrix} + \begin{bmatrix} u_{m,t} \\ u_{q,t} \end{bmatrix} \text{ (state equation)}$$

$$W_t^* \begin{bmatrix} y_{m,t} \\ y_{q,t} \end{bmatrix} = W_t^* \begin{bmatrix} I_{n_m} & 0 & 0 & 0 & 0 & 0 \\ 0 & 1/3 I_{n_q} & 0 & 1/3 I_{n_q} & 0 & 1/3 I_{n_q} \end{bmatrix} S_t \text{ (measurement equation)}$$

- Block II** (estimating the VAR parameters given Block I)

* **Time-varying intercepts** evolve as random walks: $(\Phi_{mc,t}, \Phi_{qc,t})' \equiv \Phi_{c,t} = \Phi_{c,t-1} + \nu_t$, $\nu_t \sim N(0, Q)$

* **Common stochastic volatilities** evolve as AR(1) process: $h_t = \rho h_{t-1} + \eta_t$, $\eta_t \sim N(0, \sigma_h^2)$ and resulting model is nested in u by setting $\Omega = \text{diag}(\exp(h_1), \dots, \exp(h_T))$

Priors and Hyperparameters

- Minnesota** (natural conjugate $MNIW$) priors for Φ and Ψ implemented as in ...
 - Banbura et al., 2010 (including prior on sum of coefficients)
 - Giannone et al., 2016 (priors for the long run)
- An initial state S_0 is extracted by running the SSM (in a time-constant fashion) over a pre-sample (4 years)
- Prior for $\Phi_{c,1}$ is $N(0, Q_0^{-1})$; $Q \sim IG(Q_0, k_Q \cdot Q_0 \cdot \Psi)$, where k_Q controls the amount of time variation a-priori allowed for and Q_0 corresponds to the length of the pre-sample (as in Primiceri, 2005)
- $\rho \sim N(0.9, k_\rho^2) \mathbb{I}(|\rho| < 1)$; $\sigma_h^2 \sim IG(5, k_S(5-1))$ as in Chan (2015)
- We determine suitable values for the hyperparameters $\theta = (\lambda, k_Q, k_\rho, k_S)'$ by conducting a **preliminary analysis on a grid of values**
 - Choose the grid point that led to the **best GDP forecast performance** (this is our main goal) over the last H months
 - In our application, we set $H = 48$ (i.e., 4 years), fix $k_\rho = 0.2$ as well as $k_S = 0.01$ (see Chan, 2015) and let

$$k_Q = \{\sqrt{0.001}, \sqrt{0.01}\}$$

$$\lambda = \{0.025, 0.05, 0.1, 0.25, 0.5, 0.75, 1, 5\}$$

λ controls the tightness of the Minnesota prior, i.e., the relative importance of the prior beliefs w.r.t. the information in the data

Posterior Analysis

- Conditional posterior distributions for S^T, Q, ρ, σ_h^2 and h^T are fairly standard
- We sample Ψ marginally, and then $\text{vec}(\Phi) | \Psi$:

$$\Psi | \Phi_c^T, h^T, S^T, y^T \sim IW(\hat{\Psi}, \nu_0 + T)$$
 and $\text{vec}(\Phi) | \Psi, \Phi_c^T, h^T, S^T, y^T \sim N(\text{vec}(\hat{\Phi}), K_\Phi^{-1} \otimes \Psi)$
- Conventionally drawing Φ involves (i) the inverse of the $np + 1$ square matrix K_Φ , and (ii) the Choleski decomp. of the $(np + 1)n$ square matrix $K_\Phi^{-1} \otimes \Psi$
- Using $D \sim MN(0_{m \times n}, I_m \otimes I_n) \Rightarrow C_K^{-1} DC_\Psi \sim MN(0_{m \times n}, (C_K^{-1} I_m C_K^{-1}) \otimes (C_\Psi^{-1} I_n C_\Psi^{-1}))$, we can **draw Φ more efficiently** as

$$\Phi = \hat{\Phi} + C_K^{-1} DC_\Psi$$
 where $\hat{\Phi} = (C_K C_K')^{-1} (V_{\Phi_0}^{-1} \Phi_0 + Z \Omega^{-1} X') = C_K' \backslash (C_K \backslash (V_{\Phi_0}^{-1} \Phi_0 + Z \Omega^{-1} X'))$
- For Φ_c^T , rewrite $\Phi_{c,t} = \Phi_{c,t-1} + \nu_t$ by stacking over t

$$\underbrace{\begin{bmatrix} I_n & & & \\ -I_n & I_n & & \\ & \dots & \dots & \\ & & -I_n & I_n \end{bmatrix}}_H \text{vec}(\Phi_c^T) = \text{vec}(\nu^T) [\sim N(0, I_T \otimes Q)],$$

and premultiply by H^{-1} to get the prior $\text{vec}(\Phi_c^T) | Q \sim N(0, (H'(I_T \otimes Q)^{-1} H)^{-1})$ leading to the posterior $\Phi_c^T | \Phi, \Psi, h^T, Q, S^T, y^T \sim N(\hat{\Phi}_c^T, P_{\Phi_c^{-1}})$; a similar “trick” as for K_Φ and Ψ then applies to P_{Φ_c} enabling us to **efficiently compute** $C_{P_{\Phi_c}}$ and, thereby, $\hat{\Phi}_c^T$

References

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Empirical Analysis – Data & Setup

- Three models to be compared with the benchmark MF-VAR of Schorfheide and Song (2015):
 - The “full” one including time-varying intercepts (TVI) & common stochastic volatility (CSV), labelled **TVI-MF-VAR-CSV**
 - The “intermediate” models including only one of the two features, i.e., **TVI-MF-VAR** and **MF-CSV-CSV**

Table: Data and Stylized Release Calendar

Series ID	Transf.	Pub-Lag	Description
GDP	Log	1 Qrt	Real gross domestic product, sa
INV	Log	1 Qrt	Real gross private domestic investment, sa
GOV	Log	1 Qrt	Real government consumption expenditures and gross investment, sa
UNR	1/100	1 Mth	Civilian unemployment rate, sa
HRS	Log	1 Mth	Index of aggregate weekly hours, sa
CPI	Log	1 Mth	Consumer price index for all urban consumers, sa
IPI	Log	1 Mth	Industrial production index, sa
PCE	Log	1 Mth	Personal consumption expenditures index, sa
FF	1/100	./.	Effective federal funds rate, nsa
TB	1/100	./.	10-year treasury constant maturity rate, nsa
SP500	Log	./.	S&P 500 stock index, adjusted close price, nsa

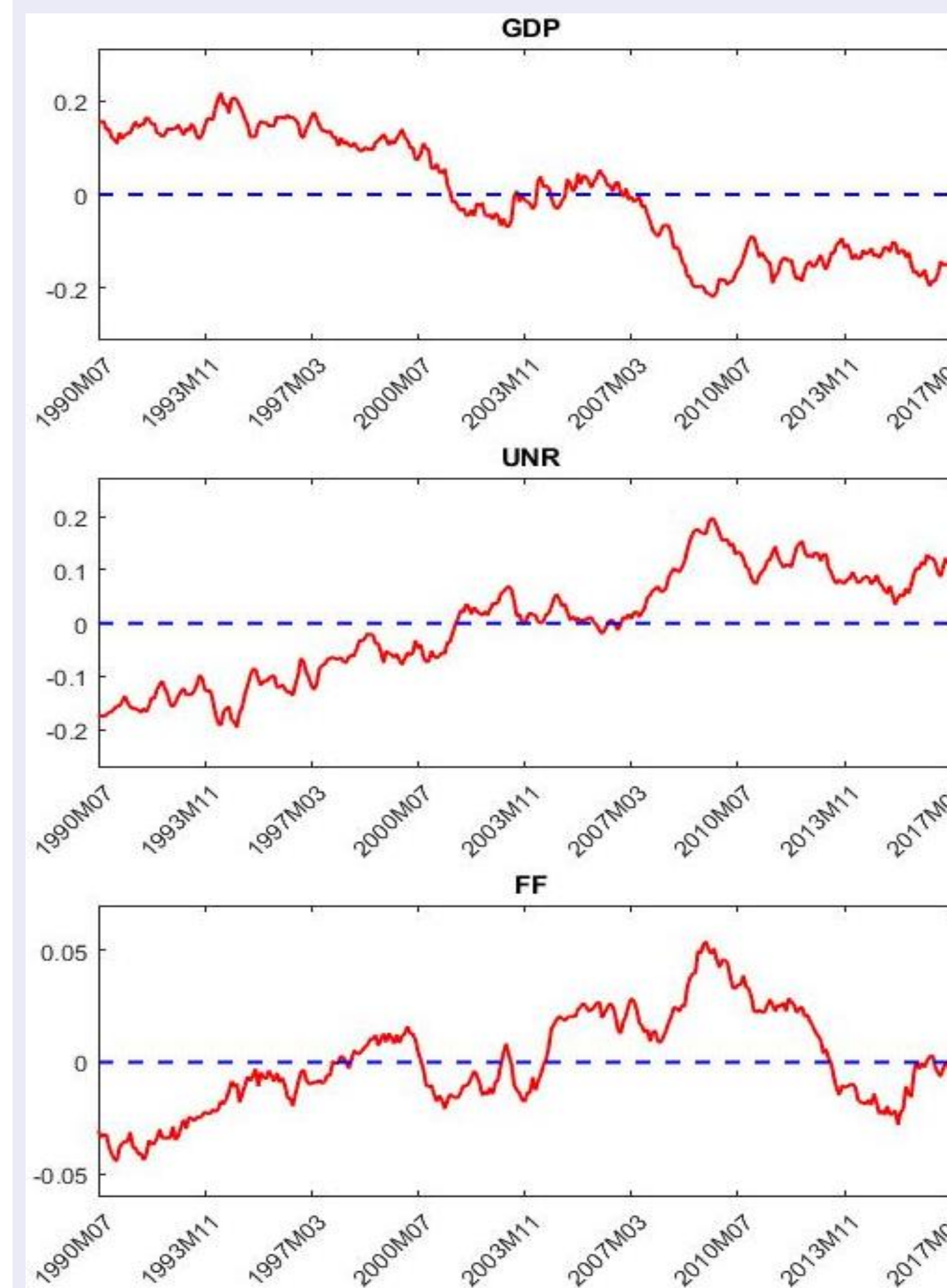
Note: GDP, INV and GOV are available on a quarterly basis, whereby second releases are considered; Data downloaded on the June 1st, 2017.

- $p = 6$ monthly lags; 1986-1989 used as pre-sample; generally, 10000 draws as burn-in, 2000 retained; forecasts assumed to be computed **at the end of each month**

Full Sample (1990:M1-2017:M7) Estimation

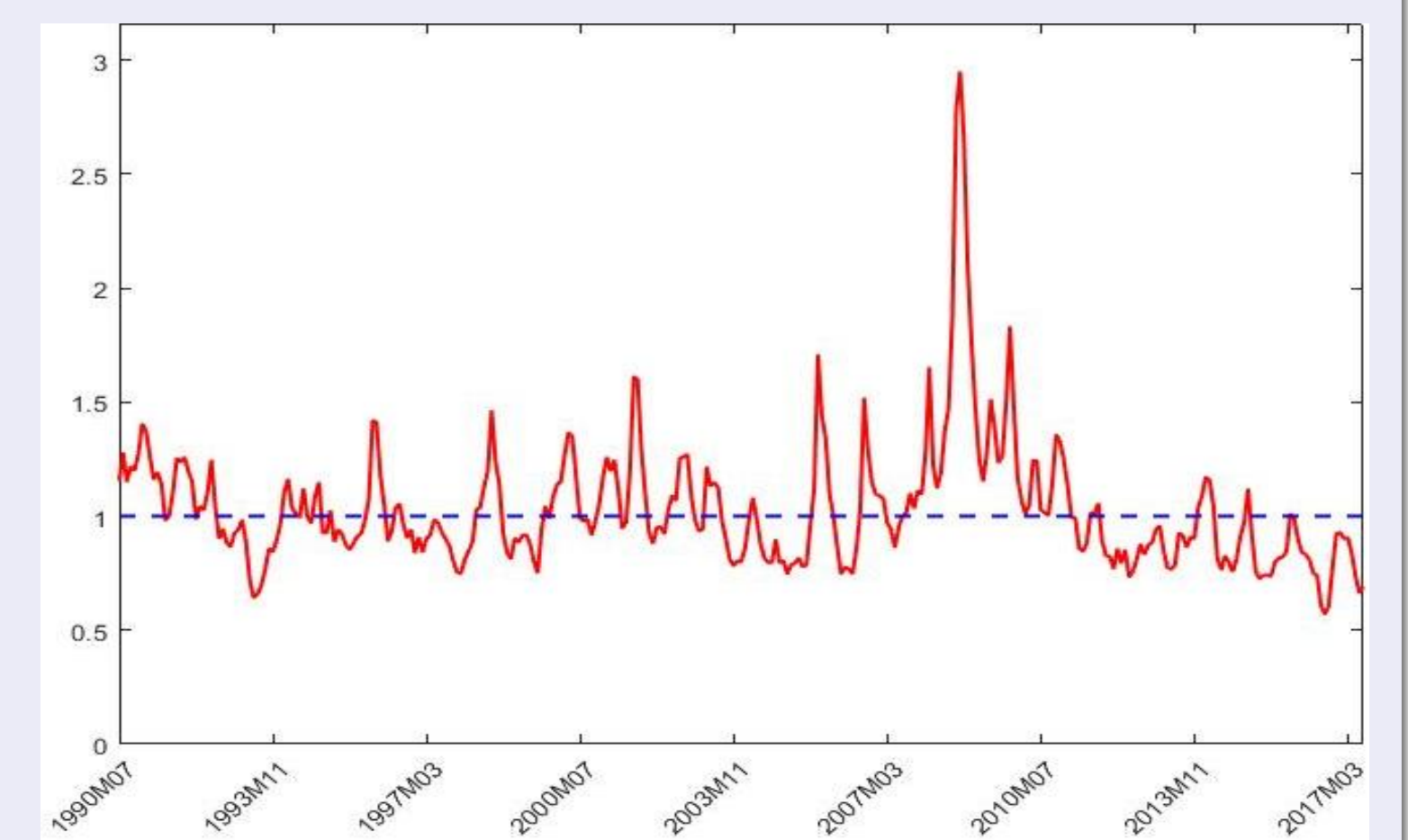
- Focus on GDP**; also consider the monthly targets **UNR** and **FF** whenever possible
- Conditional means: not fluctuations in the (long-run) growth rates; rather, **changes in the intercept’s contribution between two points in time**

Figure: Time-Varying Intercepts



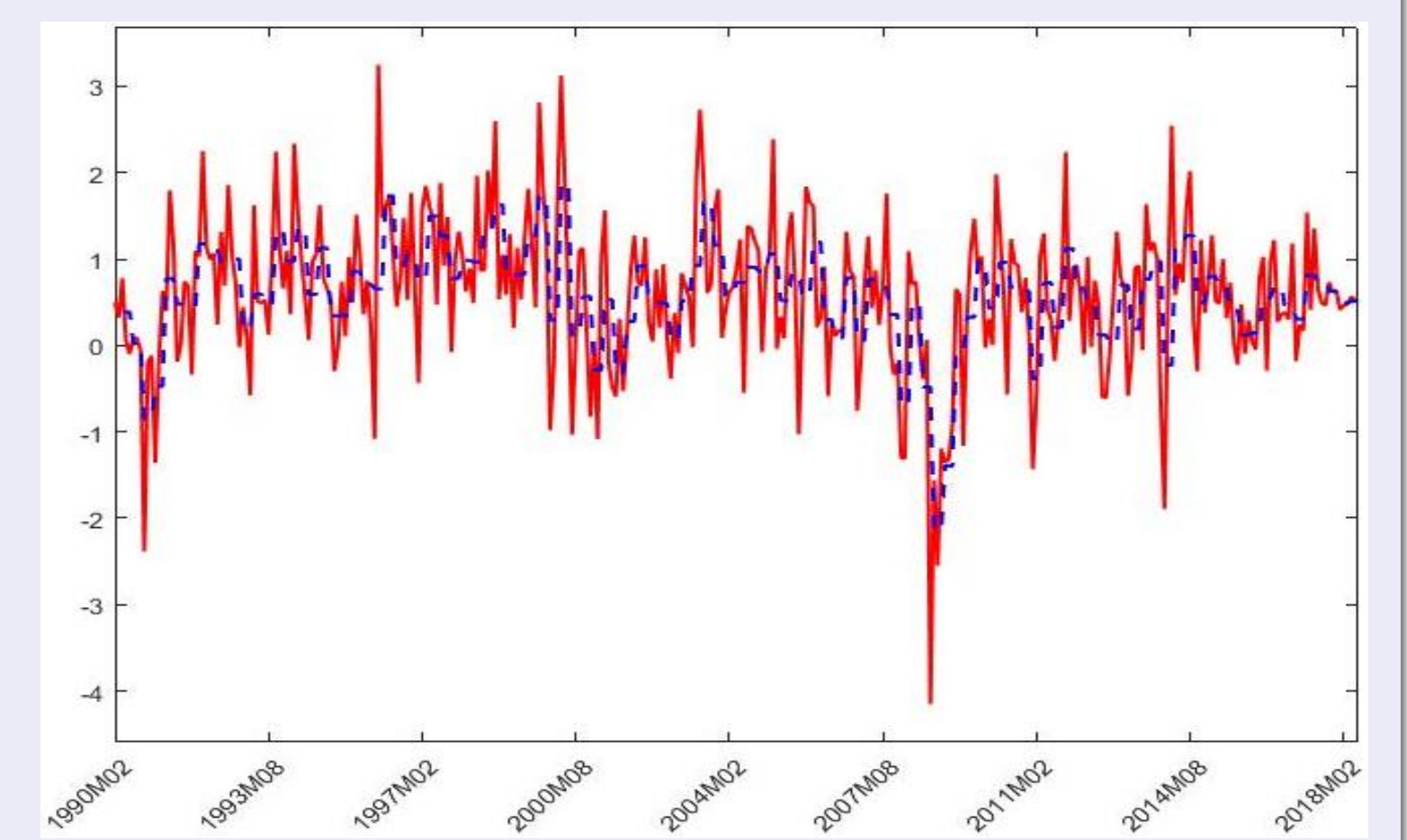
Note: Demeaned medians of the time-varying intercepts; multiplied by 100

Figure: Common Stochastic Volatility



Note: Medians of the common stochastic volatility in standard deviations

Figure: Monthly vs. Quarterly GDP Growth



Note: Growth rates of the medians of monthly GDP (including forecasts); scaled by 3 to make them comparable to quarter-on-quarter rates

- Much more movement in the monthly series due to pronounced (common) stochastic volatility
- Computation time: 41 min vs. 26min.

Forecast Evaluation Exercise

- Increasing sequence of est. samples**: 1990:M1-2006:M5, ..., 1990:M1-2017:M7; **pseudo real-time**; up to twelve forecasts for each indicator; laws of motion for TVI and CSV imposed over forecast horizon
- Forecast horizon: amount of months between a forecast is made and the end of the reference period
 - Thus, for quarterly series $h_Q = -1, 0, \dots, 10$ with $h_Q = -1$ a backcast, $h_Q = 0, 1, 2$ nowcasts and $h_Q \geq 3$ forecasts
 - For the monthly series, publication delays are shorter s.t. $h_M \geq 0$ for some series (e.g., IPI) or $h_M \geq 1$ for others (e.g., FF)
- Relative (to MF-VAR) RMSE’s and logPL’s considered to **assess point and density forecast accuracy**
- Only results with standard Minnesota priors displayed (work somewhat better than long-run priors)

Table: Forecast performance relative to the MF-VAR of Schorfheide and Song (2015)

Horizon	GDP												
	-1	0	1	2	3	4	5	6	7	8	9	10	
RMSE’s	TVI-MF-VAR	0.98	0.99	1.07	0.99	0.98	1.07	0.95	0.97	1.04	0.92	0.93	0.99
	MF-VAR-CSV	0.96	0.93	0.96	0.94	0.94	0.97	0.97	0.98	1.00	0.99	0.99	1.01
	TVI-MF-VAR-CSV	0.92	0.89	0.95	0.95	0.91	1.00	0.95	0.94	0.99	0.96	0.94	0.99
logPL’s	TVI-MF-VAR	0.06	0.15	-0.10	0.03	0.22	0.14	0.06	0.34	0.23	0.08	0.16	0.08
	MF-VAR-CSV	0.13	0.19	0.09	0.03	0.21	0.22	0.02	0.30	0.19	0.01	0.07	0.05
	TVI-MF-VAR-CSV	0.16	0.22	0.10	0.06	0.25	0.23	0.06	0.37	0.26	0.10	0.18	0.13
RMSE’s	TVI-MF-VAR	0.92	0.89	0.87	0.85	0.83	0.85	0.86	0.87	0.87	0.87	0.87	0.87
	MF-VAR-CSV	1.05	1.09	1.12	1.12	1.11	1.08	1.06	1.05	1.04	1.02	1.01	1.01
	TVI-MF-VAR-CSV	0.98	0.97	0.98	0.96	0.93	0.91	0.90	0.88	0.87	0.86	0.85	0.85
logPL’s	TVI-MF-VAR	0.07	0.10	0.15	0.24	0.35	0.50	0.80	1.17	1.24	1.29	1.29	1.21
	MF-VAR-CSV	-0.02	-0.01	0.03	0.13	0.25	0.47	0.65	0.94	0.71	0.82	0.75	0.75
	TVI-MF-VAR-CSV	0.02	0.06	0.15	0.29	0.42	0.67	0.94	1.42	1.66	1.86	1.85	1.85
RMSE’s	TVI-MF-VAR	0.94	0.95	0.95	0.96	0.96	0.95	0.95	0.95	0.95	0.96	0.97	0.97
	MF-VAR-CSV	1.01	0.99	0.99	0.98	0.97	0.96	0.96	0.96	0.96	0.96	0.97	0.97
	TVI-MF-VAR-CSV	0.95	0.93	0.93	0.93	0.92	0.91	0.91	0.91	0.91	0.92	0.93	0.93
logPL’s	TVI-MF-VAR	0.09	0.11	0.08	0.09	0.14	0.08	0.13	0.06	0.14	0.14	0.14	0.14
	MF-VAR-CSV	0.34	0.40	0.36	0.07	-0.24	-0.26	-0.39	-0.45	-0.53	-0.53	-0.50	-0.50
	TVI-MF-VAR-CSV	0.37	0.41	0.55	0.48	0.52	0.50	0.50	0.42	0.46	0.53	0.53	0.53