

Dynamics or diversity? An empirical appraisal of distinct means to measure inflation uncertainty

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Introduction

- (Inflation-) expectations play a key role in many economic models
 - Examples: New Keynesian Phillips curve, consumption smoothing, firms' investment, price setting,...
- ⇒ Under risk aversion, considering inflation uncertainty makes sense whenever inflation expectations are part of the model
- Inflation uncertainty (IU) is unobservable
 - Distinct ways to measure IU have been proposed
 - Any empirical study involving inflation risk has to motivate choice of particular uncertainty measure

Objective

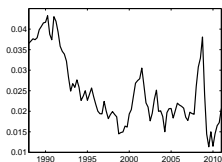
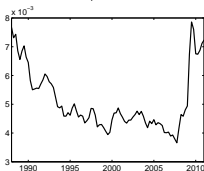
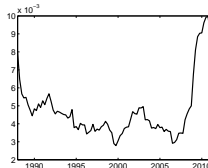
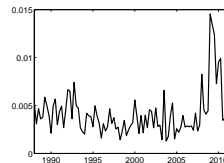
- This study: Pseudo out-of-sample forecasting 'horse race' with alternative IU measures as predictors for interest rates
- Objective: Empirical ranking of distinct approaches to measure inflation uncertainty (IU)

Distinguish two families of IU measurement:

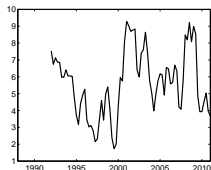
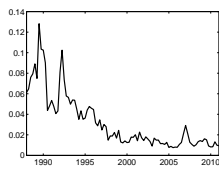
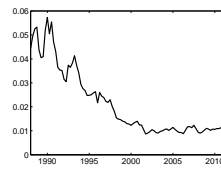
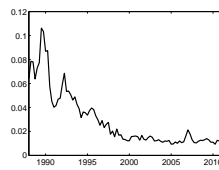
- Dynamic approaches (e.g. (G)ARCH)
- Disparity (or Dispersion) of expectations, typically based on surveys of expert forecasts, e.g. *ASA-NBER Quarterly Economic Outlook Survey*, *ZEW survey*

Median IU trajectories - 4×Dynamic (above), 4×Dispersion (below)

GARCH(1,1)

 $\sigma_{\tau+l|\tau}$  $h_{\tau+1|\tau}^{(0.1)}$  \hat{a}_τ 

ZEW-survey IU

 $\hat{\sigma}_{\tau+l|\tau}$  $\bar{\sigma}_{\tau+l|\tau}$  $\xi_{\tau+l|\tau}$ 

The figure shows the median over 18 economies. GARCH(1,1) and ZEW-survey IU are benchmark measures from the related literature

Measuring IU by means of inflation forecasting

- We consider forecast-based measures of IU
- Autoregressive (AR) scheme is among most successful models to predict inflation $\pi_t = \ln(CPI_t/CPI_{t-4})$

$$\pi_{t+l} = \alpha_0 + \alpha_1 t + \alpha_2 \pi_t + \epsilon_{t+l}, \quad t = \tau - B + 1, \dots, \tau, \quad \epsilon_{t+l} \stackrel{iid}{\sim} (0, \sigma_\epsilon^2) \quad (1)$$

- Predictions $\hat{\pi}_{\tau+l|\tau}$ obtained at forecast horizons $l \in \{1, 2, 3, 4\}$
- $\tau = T_0 - l, \dots, T - l :=$ rolling forecast origin, B is estimation window size
- time instances T_0 and T delimit period for which IU measures are obtained (1988Q1 to 2011Q1)
- Cross section comprises 18 developed economies (Austria, Belgium, Canada, Denmark, Finland, France, Germany, Ireland, Italy, Japan, Netherlands, Norway, Portugal, Spain, Sweden, Switzerland, UK, US)

Distinct ways to measure IU - 1. Dynamic measures

- 1.1 Predictive standard deviation

$$\hat{\sigma}_{\tau+\ell|\tau} = \tilde{\sigma}_\epsilon \sqrt{(1 + \mathbf{z}'_\tau (Z'_\tau Z_\tau)^{-1} \mathbf{z}_\tau)}, \quad (2)$$

with $Z_\tau :=$ design matrix of linear (AR) inflation forecasting model,
 $\mathbf{z}_\tau :=$ most recent observations for out-of-sample forecasting.

- 1.2 Exponential smoothing (Zangari 1996)

$$h_{\tau+1|\tau}^{(\lambda)} = \sqrt{\lambda(\Delta\pi_\tau)^2 + (1-\lambda)\overline{(\Delta\pi)^2}}. \quad (3)$$

In (3), $\Delta\pi_t = \pi_t - \pi_{t-1}$, and $\overline{(\Delta\pi)^2} = (1/(B-1)) \sum_{t=\tau-B+1}^{\tau-1} (\Delta\pi_t)^2$, Presetting:
 $\lambda \in \{0.1, 0.2\} \approx$ typical estimates (e.g. Bollerslev 1986)

- 1.3 Unanticipated volatility (Ball and Cecchetti 1990)

$$\hat{\alpha}_{\tau+\ell} = |\hat{\pi}_{\tau+\ell|\tau} - \pi_{\tau+\ell}|, \quad (4)$$

based on AR-implied inflation forecasts $\hat{\pi}_{\tau+\ell|\tau}$

Distinct ways to measure IU - 2. Dispersion measures

- 2.1 *Disagreement of expectations*

$$\hat{s}_{\tau+\ell|\tau} = \sqrt{(1/(J-1)) \sum_{j=1}^J (\hat{\pi}_{j,\tau+\ell|\tau} - \bar{\pi}_{\tau+\ell|\tau})^2} \quad (5)$$

from $j = 1, \dots, 5$ linear autoregressive distributed lag (ADL) forecasting models

- 2.2 *Average uncertainty* (Zarnowitz and Lambros 1987)

$$\bar{\sigma}_{\tau+\ell|\tau} = (1/J) \sum_{j=1}^J \hat{\sigma}_{j,\tau+\ell|\tau} \quad (6)$$

- 2.3 *Augmenting the disagreement measure* (cf. Lahiri and Liu 2005, Wallis 2005)

$$\xi_{\tau+\ell|\tau} = 0.5(\hat{s}_{\tau+\ell|\tau} + \bar{\sigma}_{\tau+\ell|\tau}) \quad (7)$$

- 2.4 *Alternative augmentation* (cf. Lahiri and Sheng 2010)

$$\zeta_{\tau+\ell|\tau} = 0.5(\hat{s}_{\tau+\ell|\tau} + h_{\tau+1|\tau}^{(0.1)}) \quad (8)$$

Forecasting by means of the 'augmented Fisher equation'

$$\begin{aligned}
 R_{\tau+l} = & \gamma_{10} + \gamma_{11}\tau + \sum_{p=1}^P \gamma_{12,p}\pi_{\tau-p+1} + \sum_{p=1}^P \gamma_{13,p}R_{\tau-p+1} + \\
 & + \sum_{p=1}^P \gamma_{14,p}IU_{\tau-p+l+1|\tau} + e_{\tau+l}, \quad \tau = T_0 - l, \dots, T - l \quad (9)
 \end{aligned}$$

following Levi and Makin (1979), Blejer and Eden (1979), inter alia.

- $IU_{\tau+l|\tau}$ represents a particular inflation uncertainty measure, $e_{\tau+l} \stackrel{iid}{\sim} (0, \sigma_e^2)$
- $R_{\tau+l}$: Interest rate on 10-year government bond
- Each observation $R_{\tau+l}$ from the sample period $\tau = T_0 - l, \dots, T - l$ is predicted l -steps ahead by means of a respective leave-one-out cross-validation estimate
- This yields distinct forecasts of $R_{\tau+l}$ based on alternative IU measures (2) to (8)
- Maximum lag order $P = 4 \Rightarrow 2^{12}$ distinct subset models

Subset modelling by Bayesian model averaging (BMA)

- Averaging forecasts improves predictive accuracy (Bates and Granger 1969, Timmermann 2005, Wright 2009)
- Combine forecasts from $m = 1, \dots, M = 2^{12}$ reformulations of augmented Fisher equation:

$$\hat{R}_{\tau+\ell|\tau} = \sum_{m=1}^M w_m^* \hat{R}_{\tau+\ell|\tau}^{(m)}, \quad (10)$$

•

$$w_m^* = \frac{w_m}{\sum_m w_m} \text{ and } w_m = \int L_m(\gamma^{(m)}) p_m(\gamma^{(m)}) d\gamma^{(m)}. \quad (11)$$

$L_m(\gamma^{(m)})$:= likelihood function, $p_m(\gamma^{(m)})$:= a-priori distribution of $\gamma^{(m)}$

- Based on log-likelihood $l(\gamma^{(m)}) = \ln L(\gamma^{(m)})$, posterior probabilities w_m in (11) can be approximated as

$$\ln \hat{w}_m = l(\hat{\gamma}^{(m)}) - \frac{n_m}{2} \ln(T - T_0), \quad (12)$$

$\hat{\gamma}^{(m)}$:= (Q)ML estimator of $\gamma^{(m)}$ and n_m stands for the number of right hand side variables in model m .

- Forecast combination weights obtain as w_m in (11) by $\exp\left(l(\hat{\gamma}^{(m)}) - \frac{n_m}{2} \ln(T - T_0)\right)$

Performance criterion

- Forecast ranking based on absolute forecast error (AE)

$$|e_{\tau+\ell|\tau}^{\bullet}| = |\hat{R}_{\tau+\ell|\tau}^{\bullet} - R_{\tau+\ell}| \quad (13)$$

'•' represents IU measures $\hat{\sigma}_{\tau+\ell|\tau}$, $h_{\tau+1|\tau}^{(\lambda)}$, \hat{a}_{τ} , $\hat{s}_{\tau+\ell|\tau}$, $\bar{\sigma}_{\tau+\ell|\tau}$, $\xi_{\tau+\ell|\tau}$, $\zeta_{\tau+\ell|\tau}$, $\max(\text{IU})$, $\min(\text{IU})$, $\text{median}(\text{IU})$, $\text{mean}(TS)$, $\text{mean}(DS)$.

- Frequency by which IU measure • produces forecasts among the 3 best (Stock and Watson 1999):

$$\text{TOP3}^{\bullet} = (1/((T - T_0 + 1) \times 18)) \sum_{\tau=T_0-\ell}^{T-\ell} \sum_{i=1}^{18} I(|e_{i,\tau+\ell}^{\bullet}| \leq |e_{i,\tau+\ell}^{(3)}|), \quad (14)$$

where $|e_{i,\tau+\ell}^{(3)}|$ is the 3rd smallest AE and $I(\cdot)$ is the indicator function

TOP3• frequencies

	Dynamic measures					Dispersion measures			
	$l = 1$	$l = 2$	$l = 3$	$l = 4$		$l = 1$	$l = 2$	$l = 3$	$l = 4$
$\hat{\sigma}_{\tau+l \tau}$	21.45	24.16	25.32	25.19	$\hat{s}_{\tau+l \tau}$	21.51	20.09	20.54	20.74
$h_{\tau+1 \tau}^{(0.1)}$	23.32	22.03	21.77	21.77	$\bar{\sigma}_{\tau+l \tau}$	22.35	27.65	28.62	27.97
$h_{\tau+1 \tau}^{(0.2)}$	23.26	21.77	17.57	18.09	$\varsigma_{\tau+l \tau}$	15.96	18.15	20.80	21.90
\hat{a}_{τ}	26.94	23.06	22.93	23.13	$\zeta_{\tau+l \tau}$	19.44	20.22	22.22	20.93
\overline{TS}	28.81	24.22	21.38	20.99	\overline{DS}	19.06	18.28	18.99	21.12
	Further IU statistics								
max(IU)	15.70	16.86	20.80	20.09	median	20.74	23.00	22.48	23.39
min(IU)	19.51	21.83	20.16	17.12	o	22.16	19.51	18.22	19.32

Cell entries represent the frequencies in which distinct IU measures lead to forecasts which are among the 3 most accurate ones. The row labelled as 'o' reports respective ranking frequencies for a forecasting model without an IU term.

Percentage of cases where $|e_{\tau+l|\tau}^{\bullet}| < c \times |e_{\tau+l|\tau}^{(o)}|$

	c = 1				c = 0.8			
	$\ell = 1$	$\ell = 2$	$\ell = 3$	$\ell = 4$	$\ell = 1$	$\ell = 2$	$\ell = 3$	$\ell = 4$
$\hat{\sigma}_{\tau+l \tau}$	51.03	53.29	55.49	54.84	22.22	29.07	30.75	31.20
$h_{\tau+1 \tau}^{(0.1)}$	51.87	54.20	52.71	52.00	18.09	21.25	21.25	19.77
$h_{\tau+1 \tau}^{(0.2)}$	51.74	53.94	52.84	51.74	15.50	17.64	17.70	15.31
$\hat{\alpha}_{\tau}$	49.55	51.16	52.78	52.78	26.94	29.13	28.94	28.10
$\hat{\Sigma}_{\tau+l \tau}$	51.42	53.55	54.97	54.97	26.49	31.65	35.79	34.82
$\bar{\sigma}_{\tau+l \tau}$	49.68	53.04	53.29	55.10	22.87	34.04	35.34	36.82
$\varsigma_{\tau+l \tau}$	50.19	53.10	56.07	55.56	25.32	31.65	35.47	35.92
$\zeta_{\tau+l \tau}$	50.45	52.71	54.91	54.72	27.45	33.01	37.34	35.21
max(IU)	50.45	53.10	56.20	55.62	25.26	32.11	35.47	35.59
min(IU)	49.94	53.62	52.97	50.97	18.80	22.42	22.80	22.87
median(IU)	51.55	55.88	52.26	55.62	21.45	30.62	30.30	34.30
\overline{TS}	51.16	51.36	52.39	53.23	26.94	28.42	28.55	27.78
\overline{DS}	50.65	53.23	56.14	55.62	25.97	31.91	35.85	35.79

'o' represents forecast errors for Fisher eq. WITHOUT IU term.

Comparison to benchmark measures

Percentage of cases where $|e_{\tau+\ell|\tau}^\bullet| < |e_{\tau+\ell|\tau}^{(bm)}|$

$\hat{\sigma}_{\tau+1 \tau}$	$h_{\tau+1 \tau}^{(0.1)}$	$h_{\tau+1 \tau}^{(0.2)}$	\hat{a}_τ	$\hat{s}_{\tau+1 \tau}$	$\bar{\sigma}_{\tau+1 \tau}$	$\varsigma_{\tau+1 \tau}$	$\zeta_{\tau+1 \tau}$
52.97	52.58	54.13	50.06	52.00	52.07	52.45	51.94
$\hat{\sigma}_{\tau+4 \tau}$	$h_{\tau+1 \tau}^{(0.1)}$	$h_{\tau+1 \tau}^{(0.2)}$	\hat{a}_τ	$\hat{s}_{\tau+4 \tau}$	$\bar{\sigma}_{\tau+4 \tau}$	$\varsigma_{\tau+4 \tau}$	$\zeta_{\tau+4 \tau}$
52.65	52.78	47.62	48.94	52.91	52.53	56.61	53.70

Upper panel: $bm = \text{GARCH}(1,1)$

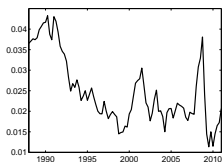
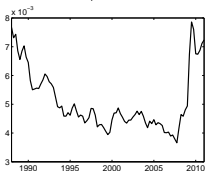
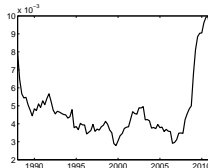
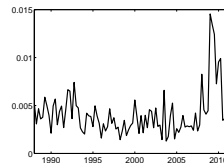
Lower panel: $bm = \text{IU based on ZEW survey}$

TOP3* for subsamples

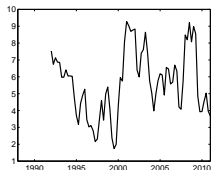
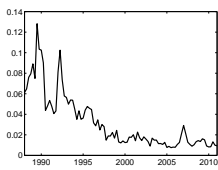
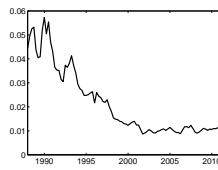
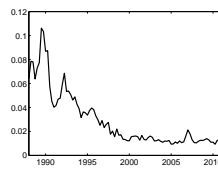
	$\ell = 1$	$\ell = 2$	$\ell = 3$	$\ell = 4$	$\ell = 1$	$\ell = 2$	$\ell = 3$	$\ell = 4$
	Turbulent periods				Calm periods			
$\hat{\sigma}_{\tau+\ell \tau}$	19.64	23.39	22.61	24.55	23.26	24.94	28.04	25.8
$\hat{\alpha}_{\tau}$	26.23	23.00	21.19	20.93	27.24	23.13	24.68	25.32
$\bar{\sigma}_{\tau+\ell \tau}$	22.61	27.91	28.29	27.00	22.09	27.39	28.94	28.94
	Sample period 1988Q1-1998Q3				Sample period 1998Q4-2011Q1			
$\hat{\sigma}_{\tau+\ell \tau}$	21.71	24.68	27.00	26.61	21.19	23.64	23.64	23.77
$\hat{\alpha}_{\tau}$	29.36	23.51	23.64	21.83	25.19	22.61	22.22	24.42
$\bar{\sigma}_{\tau+\ell \tau}$	20.67	26.61	27.91	27.26	24.03	28.37	29.33	28.68
	Higher-inflation economies				Lower-inflation economies			
$\hat{\sigma}_{\tau+\ell \tau}$	19.38	22.48	24.94	22.74	23.51	25.84	25.71	27.43
$\hat{\alpha}_{\tau}$	25.19	21.06	19.12	19.64	28.68	25.06	26.74	26.61
$\bar{\sigma}_{\tau+\ell \tau}$	20.93	27.15	29.33	28.29	23.77	27.11	27.91	27.65

Median IU trajectories - 4×Dynamic (above), 4×Dispersion (below)

GARCH(1,1)

 $\sigma_{\tau+l|\tau}$  $h_{\tau+1|\tau}^{(0.1)}$  \hat{a}_{τ} 

ZEW-survey IU

 $\hat{\sigma}_{\tau+l|\tau}$  $\bar{\sigma}_{\tau+l|\tau}$  $\xi_{\tau+l|\tau}$ 

The figure shows the median over 18 economies. GARCH(1,1) and ZEW-survey IU are benchmark measures from the related literature

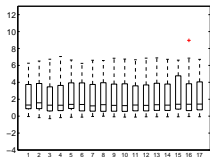
Relation between IU and $R_{\tau+\ell}$

$$\begin{aligned}
 R_{\tau+\ell} = & \gamma_{10} + \gamma_{11}\tau + \sum_{p=1}^P \gamma_{12,p}\pi_{\tau-p+1} + \sum_{p=1}^P \gamma_{13,p}R_{\tau-p+1} + \\
 & + \sum_{p=1}^P \gamma_{14,p}IU_{\tau-p+\ell+1|\tau} + e_{\tau+\ell}
 \end{aligned} \tag{15}$$

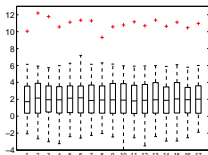
- Overall IU effect for $\tau = T_0 + 1, \dots, T$ (i.e. 1988Q1 to 2011Q1) in economies $i = 1, \dots, 18$ is denoted $\bar{\gamma}_{i\tau}^{(IU)} = \sum_{p=1}^P \gamma_{14,p}$.
- First theory: *Inflation Risk Premium*
Investors require compensation for holding non-indexed bonds (Barnea et al. 1979, Brenner and Landskroner 1983)
- Second theory: *Investment Barrier*
IU reduces demand for loanable funds since returns to real investments are more uncertain (Blejer and Eden 1979)

IU effect on $R_{\tau+l}$

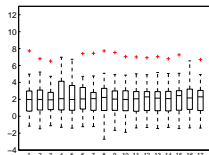
$$\hat{\sigma}_{\tau+l|\tau}$$



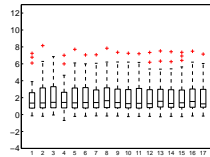
$$h_{\tau+1|\tau}^{(0.1)}$$



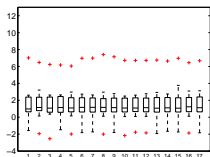
$$h_{\tau+1|\tau}^{(0.2)}$$



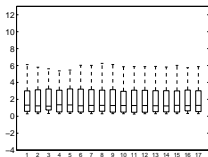
$$\hat{a}_{\tau+l}$$



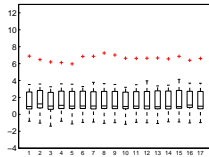
$$\hat{s}_{\tau+l|\tau}$$



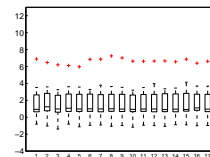
$$\bar{\sigma}_{\tau+l|\tau}$$



$$\varsigma_{\tau+l|\tau}$$



$$\zeta_{\tau+l|\tau}$$



Summary and Conclusions

We distinguish 2 families of IU measures.

- Both groups indicate IU decrease during *Great Moderation* period
- Distinct IU indication after 2008, post-Lehman

Forecast ranking shows: Dispersion outperforms Dynamic measures.

- Average over individual models' uncertainty is most informative predictor

Across time instances and economies,

impact of IU on interest is uniformly positive.

- Call IU influence on bond yields a risk premium

Measuring IU - Inflation forecasting models

Alternative ways to predict inflation:

$$\pi_{t+l} = \alpha_{10} + \alpha_{11}t + \alpha_{12}\pi_{t-1} + \alpha_{13}\tilde{y}_{t-1} + \epsilon_{t+l}, \quad t = \tau - B + 1, \dots, \tau. \quad (16)$$

$$\pi_{t+l} = \alpha_{20} + \alpha_{21}t + \alpha_{22}\pi_{t-1} + \alpha_{23}\tilde{y}_{t-1} + \alpha_{24}\bar{m}_{t-1} + \epsilon_{t+l}. \quad (17)$$

$$\pi_{t+l} = \alpha_{30} + \alpha_{31}t + \alpha_{32}\pi_{t-1} + \alpha_{33}\tilde{y}_{t-1} + \alpha_{34}\bar{m}_{t-1} + \alpha_{35}\Delta^2 oil_{t-1} + \epsilon_{t+l}. \quad (18)$$

$$\pi_{t+l} = \alpha_{40} + \alpha_{41}\tilde{\pi}_{t-1} + \epsilon_{t+l}. \quad (19)$$

- $\tilde{y}_t = y_t - \bar{y}_t$: output gap, with potential output \bar{y}_t estimated by means of HP-filter
- \bar{m}_t : core money growth
- $\Delta^2 oil_{t-1}$ oil price dynamics (WTI crude)
- $\tilde{\pi}_{t-1} = \pi_t - \bar{\pi}_t$: inflation gap $\rightarrow \tilde{\pi}_{t-1}$ in (19) resembles error-correction term